



## ORIGINAL ARTICLE

## Estimation of body fat in children with intellectual disability: development and cross-validation of a simple anthropometric method



Fabio Bertapelli <sup>a,\*</sup>, Stamatis Agiovlasitis <sup>b</sup>, Robert W. Motl <sup>c</sup>,  
Marcos M. de Barros-Filho <sup>a</sup>, Gil Guerra-Junior <sup>a</sup>

<sup>a</sup> Universidade de Campinas (Unicamp), Faculdade de Ciências Médicas, Campinas, SP, Brazil

<sup>b</sup> Mississippi State University, Department of Kinesiology, Starkville, United States

<sup>c</sup> University of Alabama at Birmingham, Department of Physical Therapy, BM, United States

Received 31 May 2021; accepted 7 January 2022

Available online 26 February 2022

### KEYWORDS

Body mass index;  
Percent body fat;  
Obesity;  
Body composition;  
Dual-energy X-Ray  
Absorptiometry;  
Childhood

### Abstract

**Objective:** Population-level monitoring of body composition requires accurate, biologically-relevant, yet feasible methods for estimating percent body fat (%BF). The aim of this study was to develop and cross-validate an equation for %BF from Body Mass Index (BMI), age, and sex among children with intellectual disability (ID). This study further aimed to examine the performance of an existing BMI-based equation (Deurenberg equation) for %BF in children with ID.

**Method:** Participants were 107 children (63 boys; aged 6-15 years) with ID randomly allocated to development ( $n = 81$ ) and cross-validation ( $n = 26$ ) samples. Dual-Energy X-Ray Absorptiometry provided the criterion %BF.

**Results:** The model including BMI, age, and sex (0 = male; 1 = female) had a significant goodness-of-fit in determining %BF ( $p < 0.001$ ;  $R^2 = 0.69$ ;  $SEE = 5.68\%$ ). The equation was:  $\%BF = -15.416 + (1.394 \times BMI) + (4.538 \times age) - (0.262 \times age^2) + (5.489 \times sex)$ . The equation was cross-validated in the separate sample based on (i) strong correlation ( $r = 0.82$ ;  $p < 0.001$ ) and non-significant differences between actual and predicted %BF ( $28.6 \pm 9.6\%$  and  $30.1 \pm 7.1\%$ , respectively); (ii) mean absolute error (MAE) = 4.4%; and (iii) reasonable %BF estimations in Bland-Altman plot (mean: 1.48%; 95% CI: 12.5, -9.6). The Deurenberg equation had a large %BF underestimation (mean: -7.1%; 95% CI: 5.3, -19.5), significant difference between actual and estimated %BF ( $28.6 \pm 9.7\%$  and  $21.5 \pm 7.0\%$ , respectively;  $p < 0.001$ ), and MAE = 8.1%.

**Conclusions:** The developed equation with BMI, sex, and age provides valid %BF estimates for facilitating population-level body fat screening among children with ID.

Study conducted at the Universidade de Campinas (Unicamp), Faculdade de Ciências Médicas, Campinas, SP, Brazil.

\* Corresponding author.

E-mail: [fbertapelli@gmail.com](mailto:fbertapelli@gmail.com) (F. Bertapelli).

<https://doi.org/10.1016/j.jpmed.2022.01.003>

0021-7557/© 2022 Sociedade Brasileira de Pediatria. Published by Elsevier Editora Ltda. This is an open access article under the CC BY-NC-ND license (<http://creativecommons.org/licenses/by-nc-nd/4.0/>).

## Introduction

Obesity is a worldwide health epidemic for individuals with and without intellectual disability (ID).<sup>1-3</sup> Children with ID, however, are nearly two times more likely to be obese than children without ID.<sup>4-6</sup> Excessive weight increases the risk for secondary health conditions such as hyperlipidemia, diabetes, respiratory and sleeping disorders among children with ID.<sup>7-9</sup> Obesity is a condition characterized by an excessive percent body fat (%BF).<sup>10</sup> Dual-Energy X-Ray Absorptiometry (DXA) is a valid method for measuring %BF in children.<sup>11,12</sup> This laboratory method, however, cannot be easily applied to population surveillance because these are time-consuming and expensive.<sup>12</sup> Among field methods, Bioelectrical Impedance Analysis (BIA) and skinfold thicknesses are common methods for measuring %BF, but it can still be inconvenient in developing countries. The use of an inexpensive and more practical tool may be of much greater immediate concern in resource-poor countries where DXA, BIA, and calipers may not be available to the general population. Another inexpensive method that requires minimal equipment involves measuring the weight and height. These measures can be easily standardized and applied within population-level body fat monitoring in developing countries where the pandemic of excessive adiposity is growing.

The Body Mass Index (BMI) has been a common method for evaluating %BF in children. Previous research indicates a moderate to strong association between BMI and %BF in children with and without ID.<sup>13-15</sup> Previous research has further developed an equation (Deurenberg Equation) for estimating %BF in individuals aged 7-15 years without disabilities.<sup>13</sup> This equation, however, may not provide accurate %BF estimates in children with ID, but this has not been examined to date. Children with ID have higher variability in anthropometric measurements compared with youth without ID.<sup>41</sup> There has further been reported some evidence of altered body composition in youth with ID, including an unproportional amount of body fat in the trunk, limbs, and whole body, compared with those without ID.<sup>16</sup> The development of a simple method may improve %BF monitoring among children with ID.<sup>17</sup> Therefore, it is important to develop and validate an equation for %BF estimation from BMI and simple demographic variables in children with ID. Moreover, once a specific equation for %BF estimation in children with ID is developed, comparing its performance with an equation for children without disabilities would further be useful for ID population-level excessive body fat surveillance among children with ID.

The aim of this study was to develop and cross-validate an equation for estimating %BF from BMI, age, and sex among children with ID. This study further aimed to examine the performance of the widely used Deurenberg BMI-based equation for estimating %BF in children with ID.

## Methods

### Participants

Participants were recruited from specialized centers for children with ID, Southeastern Brazil. The inclusion criteria were: 1) being 6-15 years old; 2) being independently ambulatory; 3) having no health conditions that could affect weight, height, or DXA measures; and 4) being able to understand the procedures. The authors included participants with health conditions that were medically treated (i.e., heart diseases, hypothyroidism, and epilepsy). Participants ( $n=110$ ) with ID aged 6-15 years volunteered for this study. The authors used this age range because past research has demonstrated a pattern of %BF development occurring in two distinct periods: %BF starts increasing linearly in girls and boys from age six; and after some fluctuations in boys during early adolescence, a similar proportional increase in %BF is observed in both girls and boys from late adolescence (about 15 years old).<sup>18</sup> The authors excluded participants who had suboptimal DXA images due to movement artifacts ( $n = 3$ ). The final sample included 107 children with ID who had all anthropometrics and DXA measures. Of those volunteers, 84 had no other disability diagnosis, 4 had cerebral palsy, 11 Down syndrome, 7 unknown genetic conditions, and 1 had microcephaly. From the total sample ( $n = 107$ ), the authors randomly selected 26 subjects (24% of the sample) to be the cross-validation sample following a previous recommendation to allocate at least 20% of the sample to cross-validation;<sup>19</sup> the remaining 81 subjects served as the sample for the equation derivation. The study was approved by the ethics committee, and all procedures were conducted according to the Declaration of Helsinki. All parents or guardians of participants provided written informed consent.

### Protocol

Participants and a parent or legal guardian attended a single session, where demographic, clinical, and anthropometric data were collected. The parent or guardian attended the session and completed a demographic and health history questionnaire about the participant. Age, ID level, sex, ethnicity, disability status, and presence of diseases were obtained from the questionnaire and from the clinical records of the specialized centers. The ID level (mild to severe ID) were obtained accordingly to the clinical records provided by the specialized center administrators. Anthropometric variables were measured by experienced technicians following standardized procedures.<sup>20</sup> Weight and height were measured in light clothes without shoes with a digital scale/stadiometer to the nearest 0.1 kg and 0.1 cm, respectively. BMI was calculated as weight in kilograms divided by height in meters squared ( $\text{kg}\cdot\text{m}^{-2}$ ). Whole-body DXA scans were performed at the iDXA (GE Healthcare Lunar, Madison, WI, USA) following the procedures outlined in the manufacturer's manual, and the %BF was determined with

the enCore™ 2011 software, version 13.6 (GE Healthcare Lunar, Madison, WI, USA). The iDXA phantom was used to ensure the high quality of DXA scans. A single experienced technician performed all DXA measurements and quality control.

### Statistical analyses

Statistical analyses were performed with SPSS Statistics 25 (IBM, Armonk, NY), and the alpha level was 0.05. The authors used independent-samples t-tests or Mann-Whitney U-test to compare continuous variables (age, anthropometrics, and %BF) and  $\chi^2$  to compare categorical variables (ethnicity, disability status, ID level, and presence of diseases) between development and cross-validation samples. The authors examined data accuracy, distributions, univariate and multivariate outliers for the dependent (%BF) and independent variables (BMI, sex, age, disability, ID level, ethnicity, and presence of diseases).<sup>19</sup> Histograms, boxplots, and Q-Q plots were used to examine normality. The authors used Cook's and Mahalanobis distances for detecting univariate or multivariate outliers.

### Prediction equation

The equation was developed using hierarchical regression models. Sequential entry of independent variables was determined following theoretical and statistical justification. The authors considered BMI, age, and sex because previous research has demonstrated a significant effect of these variables on %BF in individuals aged  $\leq 15$  years.<sup>13</sup> The authors further considered the magnitude of the bivariate associations between the dependent and independent variables. The square of BMI and age were further considered as independent variables for the examination of the nonlinear regression relationship. The authors used the Spearman rho rank-order correlation (rs) to evaluate the association between %BF and independent variables. The effect of disability status, ethnicity, presence of diseases, and ID level in the regression model were further examined. Research has demonstrated that these variables potentially predict weight gain in youth with ID.<sup>6,21,22</sup> Goodness-of-fit of the model was evaluated using  $R^2$  and Standard Error of Estimate (SEE).

### Cross-validation

The authors used the BMI-predicted %BF equation developed in the previous step to estimate the %BF in the cross-validation sample. Pearson correlation coefficient was used to examine the association between %BF determined by DXA and BMI-predicted %BF, and the square of this coefficient was calculated. A large difference between the square of Pearson's correlation coefficient in the cross-validation sample and the  $R^2$  of the regression model indicates low generalizability of the BMI-predicted %BF model.<sup>19</sup> The authors further used a paired samples t-test to examine the differences between DXA-determined and BMI-predicted %BF. Agreement between DXA-determined and BMI-predicted %BF was further evaluated with a Bland-Altman plot,<sup>23</sup> mean absolute error (MAE), and root mean square error (RMSE). Differences in MAE between disability subsamples (no other

diagnosis and other disabilities) were examined with an independent-samples t-test. The authors also estimated %BF in the cross-validation sample with a previously equation (Deurenberg Equation) for children without ID;<sup>13</sup> the equation is:

$$\begin{aligned} \%BF = & 1.51 \times BMI - 0.70 \times age - 3.6 \times sex \\ & + 1.4; \text{ BMI was in } kg \cdot m^{-2} \\ & \text{and sex (0 = female; 1 = male)}. \end{aligned} \quad (1)$$

## Results

### Samples

The development sample consisted of 81 children with ID (50 boys and 31 girls; age  $12 \pm 2.7$  years). The cross-validation sample consisted of 26 children with ID (13 boys and 13 girls; age  $11.9 \pm 2.8$  years). There were no significant differences in age, ethnicity, ID level, presence of diseases, anthropometric measures, and actual %BF between development and cross-validation samples ( $p > 0.05$ ; Table 1).

### Prediction of %BF

There were significant bivariate associations between BMI ( $rs = 0.63$ ;  $p < 0.001$ ), age ( $rs = -0.34$ ;  $p < 0.01$ ), and sex ( $rs = 0.29$ ;  $p < 0.01$ ) with %BF (Fig. 1A and B) and were sequentially entered in the hierarchical regression model. Disability, ID level, ethnicity, and presence of diseases were not significantly associated with %BF. In the regression model, BMI was a significant predictor of %BF ( $p < 0.001$ ;  $SEE = 7.70\%$  BF), explaining 41% of the variance in %BF. BMI squared was not a significant predictor and had no contribution to the variance in %BF ( $p = 0.609$ ;  $R^2$  change = 0.002). Age significantly increased the  $R^2$  from 41 to 59% and decreased the SEE from 7.70 to 6.44% BF ( $p < 0.001$ ;  $R^2$  change = 0.183). Adding age squared in the model significantly increased the  $R^2$  from 59%-63% and decreased the SEE from 6.44 to 6.22%. Sex further contributed to the model ( $p < 0.001$ ;  $R^2$  change = 0.067). Disability, ID level, ethnicity, and health condition did not contribute significantly to the final model. The final prediction model included BMI, age, age squared and sex ( $p < 0.001$ ;  $R^2 = 0.69$ ;  $SEE = 5.68\%$  BF; Table 2). The prediction equation was:

$$\begin{aligned} \%BF = & - 15.416 + (1.394 \times BMI) + (4.538 \times age) \\ & - (0.262 \times age^2) + (5.489 \times sex), \end{aligned} \quad (2)$$

BMI is in  $kg \cdot m^{-2}$  and for sex, 0 = male; 1 = female.

### Cross-validation

The authors observed a strong association between DXA-determined and BMI-predicted %BF in the cross-validation sample ( $r = 0.82$ ;  $p < 0.001$ ; Fig. 2A). The square of this correlation coefficient was nearly identical to the  $R^2$  of the prediction model (0.67 and 0.69, respectively); paired-samples t-test indicated no significant difference between DXA-

**Table 1** Demographics and anthropometrics (mean values, standard deviation, or percent) of children with intellectual disability.

Variables	Development Sample						Cross-Validation Sample					
	All (n = 81)		Boys (n = 50)		Girls (n = 31)		All (n = 26)		Boys (n = 13)		Girls (n = 13)	
	Mean	SD	Mean	SD	Mean	SD	Mean	SD	Mean	SD	Mean	SD
Age (years)	12.0	2.7	12.6	2.4	11.1 <sup>c</sup>	2.9	11.9	2.8	11.8	3.3	12.0	2.5
Weight (kg)	48.5	17.8	52.8	18.5	41.7 <sup>b</sup>	14.4	46.4	16.1	48.2	19.5	44.7	12.2
Height (cm)	150.3	16.6	155.5	15.6	142.0 <sup>a</sup>	14.8	150.6	15.1	151.6	16.8	149.6	13.9
BMI (kg·m <sup>-2</sup> )	20.9	5.1	21.3	5.3	20.2	4.9	20.0	4.7	20.3	5.7	19.8	3.7
%BF	30.7	10.0	28.4	11.1	34.2 <sup>b</sup>	6.6	28.6	9.6	24.5	10.0	32.7 <sup>c</sup>	7.5
Ethnicity												
White (%)	44.4		42.0		48.4		57.7		53.8		61.5	
Non-White (%)	55.6		58.0		51.6		42.3		46.2		38.5	
ID level												
Mild (%)	58.0		52.0		67.7		57.7		53.8		61.5	
Moderate (%)	24.7		24.0		25.8		26.9		38.5		15.4	
Severe (%)	1.2		2.0		-		7.7		7.7		7.7	
Diseases												
Yes (%)	14.8		10.0		22.6		23.1		23.1		23.1	
No (%)	79.0		86.0		67.7		73.1		76.9		69.2	

SD, standard deviation; BMI, body mass index; %BF, percent body fat determined by DXA.

<sup>a</sup>  $p \leq 0.001$ .

<sup>b</sup>  $p \leq 0.01$ .

<sup>c</sup>  $p < 0.05$  between boys and girls within the development or cross-validation sample.

There were no significant differences between development and cross-validation samples.

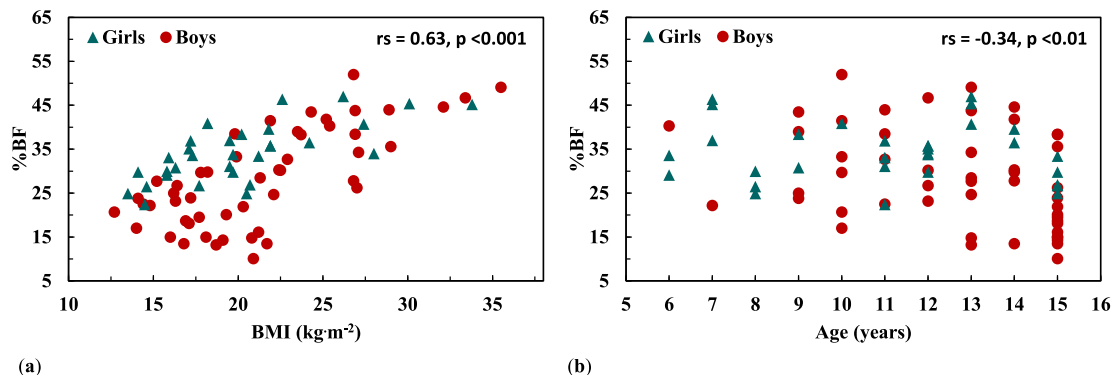
determined %BF and BMI-predicted %BF ( $28.6 \pm 9.6\%$  and  $30.1 \pm 7.1\%$ , respectively;  $p > 0.05$ ). The MAE in the cross-validation sample with the BMI-predicted %BF equation was  $4.4 \pm 3.7\%$ , and the RMSE was 6.4%. The authors further observed no significant difference in absolute errors between individuals with no other diagnosis and other disabilities ( $4.7 \pm 3.9\%$  and  $3.7 \pm 3.1\%$ , respectively,  $p > 0.05$ ). The Bland-Altman plot indicated a small mean overestimation of DXA-determined %BF and somewhat wide limits of agreement (mean error: 1.48%; 95% CI: 12.5 to -9.6; Fig. 2B). It was further observed that mean errors in children with ID with no other diagnosis and other disabilities were 1.8 and -0.2, respectively.

In contrast, the Deurenberg equation had lower %BF predictability in the cross-validation sample compared to the presently-developed equation. There was a strong

association between DXA-determined %BF and Deurenberg-predicted %BF ( $r = 0.76$ ); however, there was a large mean underestimation of DXA-determined %BF (mean error: -7.1%; 95% CI: 5.3, -19.5; Fig. 2C). Moreover, there was a significant difference between DXA-determined %BF and Deurenberg-predicted %BF ( $28.6 \pm 9.7\%$  and  $21.5 \pm 7.0\%$ , respectively;  $p < 0.001$ ), MAE = 8.1 ( $\pm 4.9\%$ ), and RMSE = 10.2%.

## Discussion

The authors developed and cross-validated an equation for estimating %BF from BMI, age, and sex in children with ID. The authors further examined the performance of the Deurenberg equation developed for children without ID. Given the ease of determining BMI, age, and sex, a population-



**Fig. 1** DXA-determined %BF as a function of BMI (a) and age (b) in children with intellectual disability.

**Table 2** Hierarchical regression models predicting DXA-determined percent body fat in children with intellectual disability.

Model	Variable	b	SE	R <sup>2</sup>	R <sup>2</sup> change	SEE
1	Intercept	4.722	3.600			
	BMI (kg·m <sup>-2</sup> ) <sup>a</sup>	1.240	0.167	0.41		7.701
2	Intercept <sup>a</sup>	21.094	4.084			
	BMI (kg·m <sup>-2</sup> ) <sup>a</sup>	1.380	0.142	0.59	0.183	6.436
	Age (years) <sup>a</sup>	-1.609	0.272			
3	Intercept	-8.085	12.129			
	BMI (kg·m <sup>-2</sup> ) <sup>a</sup>	1.362	0.137	0.63	0.032	6.221
	Age (years)	4.010	2.224			
	Age <sup>2</sup> (years) <sup>a</sup>	-0.251	0.098			
4	Intercept	-15.416	11.217			
	BMI (kg·m <sup>-2</sup> ) <sup>a</sup>	1.394	0.125			
	Age (years) <sup>a</sup>	4.538	2.034			
	Age <sup>2</sup> (years) <sup>a</sup>	-0.262	0.090			
	Sex (0 = male; 1 = female) <sup>a</sup>	5.489	1.354	0.69	0.067	5.678

b, un-standardized coefficient; SE, standard error; SEE, standard error of estimate; BMI, body mass index.

<sup>a</sup> *p* < 0.05.

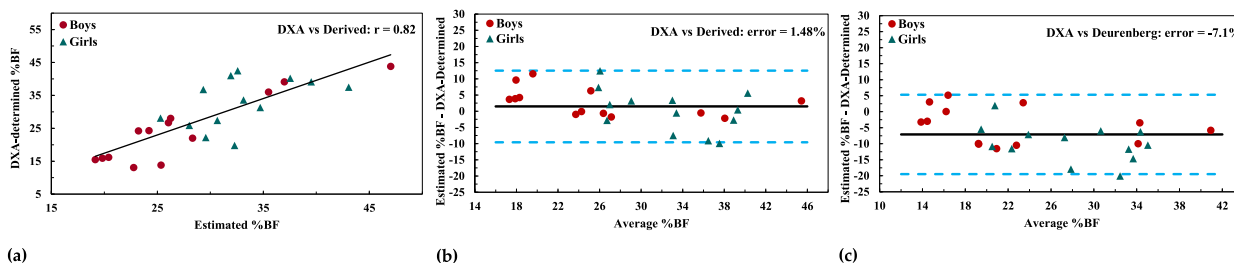
specific equation utilizing these variables may improve population-level %BF monitoring in children with ID.

BMI had a moderate linear bivariate association with %BF in the present sample of children with ID. This finding agrees with past research indicating moderate to strong associations between BMI and %BF in children with and without ID.<sup>13-15</sup> The authors’ analysis further indicated that BMI accounted for 41% of the variance in %BF, somewhat higher than the 25% of explained variance previously reported in children without ID aged 7-15 years.<sup>13</sup> Although BMI had a substantial contribution to the model, it did not explain 59% of the variance in %BF. Past research has demonstrated that %BF is substantially influenced by growth and maturation; on average, reference %BF curves indicate substantial variability in %BF from early childhood to late adolescence with higher values in girls than boys.<sup>18,24</sup> This highlights the importance of two factors in the model predicting %BF: age and sex.

Not surprisingly, age and sex, both of which had significant bivariate associations with %BF, collectively provided an additional 28% of the explained variation in %BF. This finding is in general agreement with past research that reported a significant effect of age and sex in predicting %BF with R<sup>2</sup> change of 17% in a sample of children aged 7-15 years.<sup>13</sup> The influence of age is logical since %BF changes to a significant

extent from early childhood to late adolescence.<sup>18</sup> The effect of sex is also theoretically justified because the developmental pattern of %BF is different between boys and girls; girls experience a gradual increase of %BF from early childhood to adolescence, whereas, in boys, %BF starts increasing before puberty and then gradually declines to reach its lowest point at late adolescence.<sup>18,24</sup> Although it is well-known that %BF changes substantially during growth and maturation among youth without ID, little information is available about the developmental patterns of %BF in children with ID. Recent studies showed that %BF in girls with Down syndrome was somewhat higher than in boys.<sup>15</sup> Apart from sex, %BF seems to vary as a function of age in children with Down syndrome.<sup>25</sup> In summary, BMI, age, and sex accounted for a substantial portion of the variance in %BF, and this is consistent with past experimental evidence. The authors believe that the next step involves examining easily-determined factors that could potentially account for the 31% of unexplained variance in %BF and improve prediction among youth with ID. Such factors may include the stage of biological maturation.<sup>26</sup>

The BMI-predicted %BF equation the authors developed had the reasonable predictive ability to estimate %BF in a separate sample of children with ID. The small difference



**Fig. 2** DXA-%BF against the derived Equation-%BF in children with intellectual disability in the cross-validation sample (a). Bland-Altman plots of the difference between the %BF measurements (y-axis) against their average (x-axis) using the presently BMI-predicted %BF equation (b) and the Deurenberg equation (c) in boys and girls with intellectual disability; solid and dotted lines show mean and 95% limits of agreement, respectively.

between the  $R^2$  of the BMI-predicted %BF equation and the square of the correlation between the estimated and DXA-determined %BF in the cross-validation sample indicated the high generalizability of the equation. Moreover, the authors did not find any significant difference between DXA- and BMI-predicted %BF mean values and relatively small absolute and root mean square error. This finding is supported by previous research predicting %BF from BMI, age, and sex in the general population;<sup>13</sup> using their equation in a separate sample of children, these researchers also found a relatively small difference between predicted and observed %BF. When the authors examined the study's Bland-Altman plot, we found a small overestimation of DXA-determined %BF and somewhat wide 95% limits of agreement. This may indicate the possibility of a small overall bias but lower predictability when the equation is applied to individual children with ID rather than at a group level.

The present study's BMI-predicted equation demonstrated higher agreement for %BF estimation compared with the previous equation developed for children without ID, as indicated by a stronger association and smaller prediction errors of the equation against the DXA-%BF criterion method. One biological explanation for the differences between the equations is that children with ID may have higher variability in total and regional body fat and fat-free mass distributions than children without ID.<sup>16</sup> Another consideration is that the equation by Deurenberg et al. was derived from a sample of leaner children than the present participants with ID. For example, the mean %BF for the Deurenberg sample was 18%, compared to the study's sample mean of 30%. It is also important to note that the Deurenberg equation was developed and cross-validated using a different laboratory reference criterion for %BF determination (underwater weighing) which results in markedly different %BF estimates compared to DXA in the general population of youth.<sup>27,28</sup> Lastly, it should be considered that the Deurenberg equation included age in the model, and the equation included age and age squared. It is evident that an equation for predicting %BF specifically for children with ID is more appropriate than a generalized equation developed for the general population.

The present BMI-predicted %BF equation has clinical and research implications. Researchers may use the present equation to estimate population-level %BF in children with ID where the laboratory method is not available. This is particularly relevant for advancing large-scale epidemiological surveillance research tracking body fat from four simple measurements (i.e., weight, height, age, and sex). Children with ID are at higher risk for obesity than children without ID.<sup>4-6</sup> However, because the present prediction equation demonstrated somewhat wide limits of agreement, caution should be exercised when using the equation for clinical practice. Researchers and health providers are cautious when applying the present equation to subpopulations of children with ID, such as those with cerebral palsy or Down syndrome. Past researchers have reported that these individuals may demonstrate higher regional variability in anthropometric measurements compared with youth with ID.<sup>29</sup> For instance, youth with Down syndrome are at higher risk for being overweight than youth with ID, whereas youth

with cerebral palsy are at higher risk for being underweight compared to other disabilities.<sup>29</sup>

There are some limitations that should be considered. First, the convenience samples may not be representative of the whole population of youth with ID. Second, the authors included a specific age range (6-15 years), and the results may not generalize to youth outside this range. Lastly, although the authors observed that, in youth with other disabilities, differences between techniques were nearly zero in the Bland-Altman, small sample sizes might limit the generalizability of the equation to these sub-populations. The following strengths should also be considered. The presently %BF equation was developed from DXA. This equation was developed from simple measures and was cross-validated on a separate sample.

In conclusion, BMI, age, and sex were significant predictors of %BF in children with ID. An equation estimating %BF from these predictors in children with ID was relatively accurate. The results have implications for improving public health surveillance, facilitating population-level body fat screening, and advancing excessive %BF prevention research among children with ID.

## Funding

This work was supported by the Sao Paulo Research Foundation (FAPESP grants: 2019/07103-6; 2017/13071-4; 2018/02677-1).

## Conflicts of interest

The authors declare no conflicts of interest.

## References

1. NCD Risk Factor Collaboration (NCD-RisC). Worldwide trends in body-mass index, underweight, overweight, and obesity from 1975 to 2016: a pooled analysis of 2416 population-based measurement studies in 128•9 million children, adolescents, and adults. *Lancet*. 2017;390:2627–42.
2. Sadowsky M, McConkey R, Shellard A. Obesity in youth and adults with intellectual disability in Europe and Eurasia. *J Appl Res Intellect Disabil*. 2020;33:321–6.
3. Stancliffe RJ, Lakin KC, Larson S, Engler J, Bershadsky J, Taub S, et al. Overweight and obesity among adults with intellectual disabilities who use intellectual disability/developmental disability services in 20 USS States. *Am J Intellect Dev Disabil*. 2011;116:401–18.
4. Maïano C, Hue O, Morin AJ, Moullec G. Prevalence of overweight and obesity among children and adolescents with intellectual disabilities: a systematic review and meta-analysis. *Obes Rev*. 2016;17:599–611.
5. Emerson E, Robertson J, Baines S, Hatton C. Obesity in British children with and without intellectual disability: cohort study. *BMC Public Health*. 2016;16:644.
6. Segal M, Eliasziw M, Phillips S, Bandini L, Curtin C, Kral TV, et al. Intellectual disability is associated with increased risk for obesity in a nationally representative sample of USS children. *Disabil Health J*. 2016;9:392–8.
7. Rimmer JH, Yamaki K, Davis BM, Wang E, Vogel LC. Obesity and overweight prevalence among adolescents with disabilities. *Prev Chronic Dis*. 2011;8. A41Epub 2011 Feb 15. PMID: 21324255; PMCID: PMC3073434.

8. Phillips KL, Schieve LA, Visser S, Boulet S, Sharma AJ, Kogan MD, et al. Prevalence and Impact of unhealthy weight in a national sample of US adolescents with autism and other learning and behavioral disabilities. *Matern Child Health J.* 2014;18:1964–75.
9. Vanhelst J, Bui-Xuan G, Fardy PS, Mikulovic J. Relationship between sleep habits, anthropometric characteristics and lifestyle habits in adolescents with intellectual disabilities. *Res Dev Disabil.* 2013;34:2614–20.
10. World Health Organization (WHO). *Obesity: Preventing and Managing the Global Epidemic.* Geneva: WHO; 2000, Technical Report Series.
11. Lohman TG, Milliken LA. *ACSM's Body Composition Assessment.* Champaign: Human Kinetics. 2020.
12. Heymsfield SB, Lohman TG, Wang Z, Going SB. *Human body composition.* Champaign: Human Kinetics; 2005. p. p 15, 2nd edn. 200515–33.
13. Deurenberg P, Weststrate JA, Seidell JC. Body mass index as a measure of body fatness: age- and sex-specific prediction formulas. *Br J Nutr.* 1991;65:105–14.
14. Lohman TG, Caballero B, Himes JH, Davis CE, Stewart D, Houtkooper L, et al. Estimation of body fat from anthropometry and bioelectrical impedance in Native American children. *Int J Obes.* 2000;24:982–8.
15. González-Agüero A, Matute-Llorente Á, Gómez-Cabello A, Vicente-Rodríguez G, Casajús JA. Percentage of body fat in adolescents with Down syndrome: estimation from skinfolds. *Disabil Health J.* 2017;10:100–4.
16. González-Agüero A, Ara I, Moreno LA, Vicente-Rodríguez G, Casajús JA. Fat and lean masses in youths with Down syndrome: Gender differences. *Res Dev Disabil.* 2011;32:1685–93.
17. Matson JL, Matson ML. Comorbid conditions in individuals with intellectual disabilities. In: Matson JL, ed. *Autism Child Psychopathol.* New York: Springer; 2015:275–98.
18. Malina RM, Bouchard C, Bar-Or O. *Growth, maturation, and physical activity.* Champaign: Human Kinetics; 2004. p. p 101, 2nd edn. 2004101–19.
19. Tabachnick BG, Fidell LS. *Using Multivariate Statistics.* 7th edn. Boston: Pearson; 2019.
20. Lohman TG, Roche AF, Martorell R. *Anthropometric Standardization Reference Manual.* Champaign, IL: Human Kinetics; 1988. p. 31–58.
21. van Gameren-Oosterom HB, van Dommelen P, Schönbeck Y, Oudesluys-Murphy AM, van Wouwe JP, Buitendijk SE. Prevalence of overweight in Dutch children with down syndrome. *Pediatrics.* 2012;130:e1520–6.
22. Hellings JA, Zarccone JR, Crandall K, Wallace D, Schroeder SR. Weight gain in a controlled study of risperidone in children, adolescents and adults with mental retardation and autism. *J Child Adolesc Psychopharmacol.* 2001;11:229–38.
23. Bland JM, Altman DG. Measuring agreement in method comparison studies. *Stat Methods Med Res.* 1999;8:135–60.
24. Ogden CL, Li Y, Freedman DS, Borrud LG, Flegal KM. Smoothed percentage body fat percentiles for USS children and adolescents, 1999-2004. *Natl Health Stat Report.* 2011.
25. Pitchford EA, Adkins C, Hasson RE, Hornyak JE, Ulrich DA. Association between physical activity and adiposity in adolescents with down syndrome. *Med Sci Sports Exerc.* 2018;50:667–74.
26. Taylor RW, Grant AM, Williams SM, Goulding A. Sex differences in regional body fat distribution from pre- to postpuberty. *Obesity.* 2010;18:1410–6.
27. Wells JC, Williams JE, Chomtho S, Darch T, Grijalva-Eternod C, Kennedy K, et al. Body-composition reference data for simple and reference techniques and a 4-component model: A new UK reference child. *Am J Clin Nutr.* 2012;96:1316–26.
28. Fields DA, Goran MI. Body composition techniques and the four-compartment model in children. *J Appl Physiol.* 2000;89: 613–20.
29. Rimmer JH, Yamaki K, Lowry DM, Wang E, Vogel LC. Obesity and obesity-related secondary conditions in adolescents with intellectual/developmental disabilities. *J Intellect Disabil Res.* 2010;54:787–94.